

Isotropy

Remark. For a random vector $X = (x_1, \dots, x_n)$, recall that

$$\begin{aligned}\text{cov}(X) &= \mathbb{E} X \cdot X^T - \mu \cdot \mu^T \\ &= \Sigma - \mu \cdot \mu^T\end{aligned}$$

The (i, j) -th entries of $\text{cov}(X)$ and Σ are

$$\text{cov}(X)_{ij} = \mathbb{E} (X_i - \mathbb{E} X_i)(X_j - \mathbb{E} X_j)$$

\uparrow [covariance of X_i and X_j]

and

$$\Sigma_{ij} = \mathbb{E} X_i X_j.$$

Def. A random vector $X = (x_1, \dots, x_n)$ is called **isotropic** if

$$\Sigma(X) = \mathbb{E} X X^T = I_n,$$

where I_n is the $n \times n$ identity matrix.

Isotropy is the higher-dimensional analogue of unit variance (if mean zero) or $\mathbb{E} X^2 = 1$.

For random variables, it's common to center and scale to get the **standard score** (or **z-score**)

$$Z = \frac{X - \mathbb{E}X}{\sqrt{\text{Var}(X)}} \quad \left(\begin{array}{l} \text{mean zero} \\ \text{unit variance} \end{array} \right)$$

We can do something similar for random vectors:

Reduction to Isotropy

Suppose X has mean μ and $\text{cov}(X) = \Sigma$ is invertible. Then

$$Z = \Sigma^{-1/2} \cdot (X - \mu)$$

is isotropic with mean zero.

Characterizing Isotropy

Lemma 3.2.3 A random vector X in \mathbb{R}^n is isotropic if and only if

$$\mathbb{E}[\langle X, x \rangle^2] = \|x\|_2^2$$

for all $x \in \mathbb{R}^n$.

Proof. Two real symmetric matrices, A and B , are equal if and only if

$$x^T A x = x^T B x$$

for all $x \in \mathbb{R}^n$. Why?

[Hint: consider $x^T(A-B)x = 0$ for certain choices of x and use symmetry.]

Recall that X is isotropic when $\mathbb{E}XX^T = I_n$.

Hence X is isotropic if and only if

$$x^T(\mathbb{E}XX^T)x = x^T \cdot I_n \cdot x \text{ for all } x \in \mathbb{R}^n.$$

The LHS is

$$\begin{aligned}\mathbb{E}(x^T \underline{X} \underline{X}^T x) &= \mathbb{E} \langle x, \underline{X} \rangle \langle \underline{X}, x \rangle \\ &= \mathbb{E} \langle \underline{X}, x \rangle^2\end{aligned}$$

On the other hand, the RHS is

$$x^T \cdot I_n \cdot x = x^T \cdot x = \langle x, x \rangle = \|x\|_2^2. \quad \square$$

Equivalently, \underline{X} is isotropic if and only if

$$\mathbb{E} \langle \underline{X}, u \rangle^2 = 1$$

for all unit vectors $u \in \mathbb{R}^n$.

Recall that $\langle \underline{X}, u \rangle$ is the projection of \underline{X} along u . [a one-dimensional marginal of \underline{X}] Hence, the lemma says that isotropic distributions tend to extend evenly in all directions.

Lemma 3.2.4 Let X (in \mathbb{R}^n) be isotropic,
Then

$$\mathbb{E} \|X\|_2^2 = n.$$

If X and Y are independent and isotropic, then

$$\mathbb{E} \langle X, Y \rangle^2 = n.$$

Proof. We have

$$\mathbb{E} \|X\|_2^2 = \mathbb{E} \langle X, X \rangle$$

$$= \mathbb{E} X^T \cdot X$$

trace \swarrow

$$= \mathbb{E} \text{tr}(X^T \cdot X)$$

$$[\text{tr}(AB) = \text{tr}(BA)] \quad = \mathbb{E} \text{tr}(X \cdot X^T)$$

$$[\text{linearity of tr}] \quad = \text{tr}(\mathbb{E} X \cdot X^T).$$

Since X is isotropic, this is $\text{tr}(I_n) = n.$

For the second part, we condition on Y . The law of total expectation says that

$$\mathbb{E} \langle X, Y \rangle^2 = \mathbb{E}_Y \mathbb{E}_X [\langle X, Y \rangle | Y].$$

$$[\text{Lemma 3.2.3}] = \mathbb{E}_Y \|Y\|_2^2$$

$$[\text{first part of proof}] = n.$$

□

Remark. We've now seen that

$$\|X\|_2^2, \|Y\|_2^2 \approx n$$

and also

$$\sqrt{n} \approx |\langle X, Y \rangle|$$

$$= \|X\|_2 \cdot \|Y\|_2 |\cos \theta| \quad \text{[angle between } X, Y \text{]}$$

$$\approx n \cdot |\cos \theta|.$$

Then $|\cos \theta| \approx \frac{1}{\sqrt{n}}$, which is ≈ 0 for large n .
Hence $\theta \approx \pm \frac{\pi}{2}$. That is, independent isotropic random vectors tend to be nearly orthogonal!

Examples of Isotropic Distributions

1) If X_1, X_2, \dots, X_n are independent random variables with mean zero and unit variance, then $X = (X_1, \dots, X_n)$ is isotropic. Why?
The entries of Σ are

$$\begin{aligned} \mathbb{E} X_i X_j &= \begin{cases} \mathbb{E} X_i^2 & \text{if } i=j, \\ \mathbb{E} X_i X_j & \text{if } i \neq j, \end{cases} \\ &= \begin{cases} 1 & \text{if } i=j, \\ 0 & \text{if } i \neq j. \end{cases} \end{aligned}$$

Thus $\Sigma = I_n$, so X is isotropic.

2) (special case) If each X_i in 1) is a symmetric Bernoulli random variable [so $P(X_i = 1) = P(X_i = -1) = \frac{1}{2}$], then X is a symmetric Bernoulli random vector. Equivalently, X is the uniform distribution on the unit discrete cube $\{-1, 1\}^n$ in \mathbb{R}^n .

Coordinates need not be independent, however.

3) Suppose X is uniformly distributed on the sphere of radius \sqrt{n} centered at the origin in \mathbb{R}^n . If $X = (X_1, \dots, X_n)$, then

$$X_1^2 + \dots + X_n^2 = n.$$

Thus
$$X_n = \pm \sqrt{X_1^2 + \dots + X_{n-1}^2},$$

which depends on the other coordinates.

[Also called the spherical distribution in \mathbb{R}^n .]

Multivariate Gaussians

A random vector $g = (g_1, \dots, g_n)$ is said to have the **standard normal** distribution, written $g \sim N(0, I_n)$, if $g_1, \dots, g_n \sim N(0, 1)$ are independent.

Since the coordinates are independent with mean zero and unit variance, g is isotropic.

Its PDF is given by

$$f(x) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}} e^{-x_i^2/2}$$

$$= \frac{1}{(2\pi)^{n/2}} e^{-\|x\|_2^2/2}, \quad x \in \mathbb{R}^n.$$

Note. The density only relies on the length of x , so it is rotation invariant.

Prop. Suppose U is an orthogonal $n \times n$ matrix, and $g \sim N(0, I_n)$. Then

$$Ug \sim N(0, I_n).$$

We get a general normal distribution as follows. Let $\mu \in \mathbb{R}^n$ and suppose Σ is an invertible, positive-semidefinite matrix. If $Z \sim N(0, I_n)$, then

$$X := \mu + \Sigma^{1/2} Z$$

is normally distributed with mean μ and covariance matrix Σ (check!). Its density is given by

$$f_X(x) = \frac{1}{(2\pi)^{n/2} (\det \Sigma)^{1/2}} e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1} (x-\mu)}$$

for $x \in \mathbb{R}^n$.

Remark. If $X \sim N(\mu, \Sigma)$, then the coordinates of X are independent if and only if they are uncorrelated! In such case, it follows that $\Sigma = I_n$ (check!).

[Generally, independence implies uncorrelated, but not the other way around.]

Also, we know from **Theorem 3.1.1** that $g \sim N(0, I_n)$ concentrates around the sphere of radius \sqrt{n} . Thus, in high dimensions, the standard normal is close to the uniform distribution on the sphere. This behavior differs greatly from the standard normal in \mathbb{R} !